

Ζιόναρδος Μητροφίεριος:

1/7/2024

Θεοφάνης Καραϊρής, Ζυγοπέντες Βαύλος και Τουκέη,
Thresholds για ηραρχικώς νοτία μέσο.

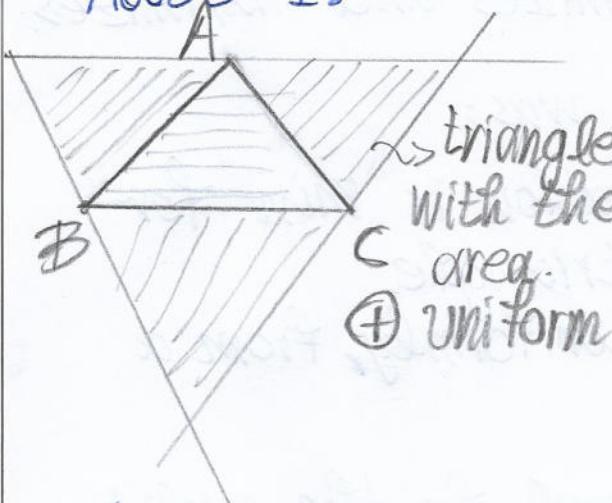
• 1864: Educational Times.

Sylvester: What is the chance of getting a convex quadrilateral after taking 4 points from the infinite plane?
→ 6 different answers to the problem:

1. Sylvester and Cayley $\approx \frac{3}{4}$.

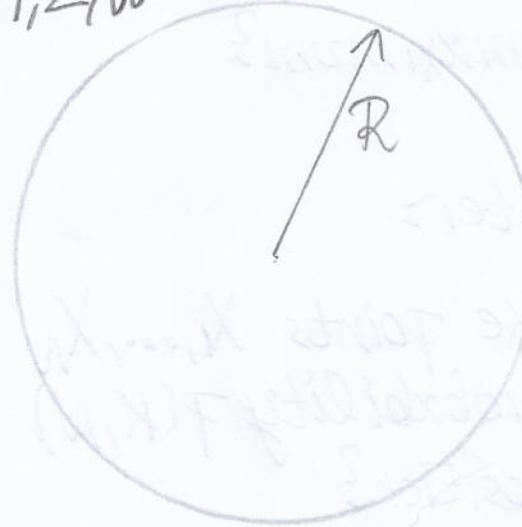
2. Woolhouse $\approx 1 - \frac{35}{12\pi^2} \approx 0.404$.

~ About 1:



~ About 2:

X, Y, Z, W :



• Introduce the random variables

I_x, I_y, I_z, I_w as:

• $I_x := \begin{cases} 1, & \text{if } X \in \text{inside } YZW \\ 0, & \text{otherwise} \end{cases}$

Let $I = I_x + I_y + I_z + I_w$ (*).

• We want the probability of
 $\{I \geq 0\} (= P)$

$$\mathbb{E}[I] = 0 \cdot p + 1 \cdot (1-p) = 1-p.$$

But from (*) $\mathbb{E}[I] = \mathbb{E}[I_x + I_y + I_z + I_w] =$ ①

$$= 4 \cdot \mathbb{E}[I_x]$$

And $\mathbb{E}[I_x] = \mathbb{E}[\mathbb{E}[I_x | Y, Z, W]] = \mathbb{E}[\text{P}[I_x | Y, Z, W]]$
 $= \mathbb{E}\left[\frac{\text{area}(YZW)}{\text{area}(D)}\right]$ and the radius R doesn't play
any role so we can take $R \rightarrow \infty$.

*: After the 6 answers it was concluded that the problem wasn't well posed. So:

- If we take our points from a convex set, what is the maximum and the minimum probability?
- This is equivalent to find the convex set K for which $\mathbb{E}\left[\frac{\text{area}(YZW)}{\text{area}(K)}\right]$ minimizes and maximizes.

The first to answer the question was:

Blaschke (1928): The above expectation is min for the disc and maximum for the triangle.

Indeed, if we take X_1, X_2, \dots, X_N uniformly from a convex set K , with $|K|=1$, then:

$\mathbb{E}[\text{conv}\{X_1, X_2, \dots, X_N\}^c]$ is minimal for the euclidean ball.

Groemer (1973): What about the maximum?

Open, even for $n=3$.

The general problem of Sylvester:

Take $K \subseteq \mathbb{R}^n$ convex body and take points X_1, \dots, X_N uniformly from K . What is the probability $P(K, N)$ that the points are in convex position?

~> Exact computation is only known for the $K = \mathbb{B}_2^n$ for N , $P(\Delta^n, n+2)$ is also known.

About the behaviour:

Theorem (Barany, Füredi 1988):

If $K = \mathbb{B}_2^n$ and $N < \frac{2^{n/2}}{n}$ then $P(\mathbb{B}_2^n, N) \xrightarrow{n \rightarrow \infty} 0$.

and $N > n \cdot 2^{n/2}$ then $P(\mathbb{B}_2^n, N) \xrightarrow{n \rightarrow \infty} 1$.

▷ Analogous Question: If X_1, X_2, \dots, X_N uniformly from $K \subseteq \mathbb{R}^n$ and $K_N := \text{conv}\{X_1, \dots, X_N\}$

$\frac{|E[K_N]|}{|K|} \in [0, 1]$. So we ask for the number of points N for which $\frac{|E[K_N]|}{|K|} \xrightarrow{n \rightarrow \infty} 0$, $\frac{|E[K_N]|}{|K|} \xrightarrow{n \rightarrow \infty} 1$.

The first result (Dyer, Füredi, McDiarmid '92):

If $K = \mathbb{B}_{\infty}^n = [-1, 1]^n$ and fix $\epsilon \in (0, 1)$ and $K := \log \frac{2}{\sqrt{e} \cdot e^{n/2}}$. Then if $N > e^{Kn(1+\epsilon)}$, then:

$\frac{|E[K_N]|}{|[-1, 1]^n|} \xrightarrow{n \rightarrow \infty} 1$ and if $N < e^{Kn(1-\epsilon)}$, then:

$\frac{|E[K_N]|}{|[-1, 1]^n|} \xrightarrow{n \rightarrow \infty} 0$

The second Result [After Break]

If $K = \mathbb{B}_2^n$, then fix $\epsilon \in (0, 1)$. Then, if $N > e^{\frac{(1+\epsilon)}{2} \cdot \log n}$

then $\frac{\mathbb{E}[|Kn|]}{|K|} \rightarrow 1$ and if $N < e^{\frac{(1-\epsilon)}{2} \cdot \log n}$

We have: $\frac{\mathbb{E}[|Kn|]}{|K|} \xrightarrow[n \rightarrow +\infty]{} 0$

Giannopoulos, Patis, B.: In the case of \mathbb{B}_2^n we can take
 $\epsilon = \epsilon(n) = \frac{1}{\log(n)}$

2020: Tkocz, Frieze, Pedgen. The case of the simplex.
There are constants c_1 and c_2 s.t. if $N > e^{c_1 \cdot n}$, then:

$$\frac{\mathbb{E}[|Kn|]}{|D^n|} \xrightarrow[n \rightarrow +\infty]{} 1 \text{ and if } N < e^{c_2 \cdot n} \text{ then}$$

$$\frac{\mathbb{E}[|Kn|]}{|D^n|} \xrightarrow[n \rightarrow +\infty]{} 0$$

Generalizations: For a random vector X in \mathbb{R}^n , we define $\Lambda(\xi) := \log \mathbb{E}[e^{\langle \xi, X \rangle}]$ (log-Laplace transformation)

- Λ : convex function $\Rightarrow \exists$ dual of Λ .

- For $x \in \mathbb{R}^n$, we define: $\Lambda_X^*(x) := \sup_{\xi \in \mathbb{R}^n} (\langle x, \xi \rangle - \Lambda_X(\xi))$

[Cramer Transformation \rightsquigarrow Large Deviation Theory].

- If $\{X_i\}_{i=1}^n$ are r.v then $\forall \delta > 0$

$$- \mathbb{P}\left[\left|\frac{\sum X_i}{n}\right| \geq \delta\right] \xrightarrow[n \rightarrow +\infty]{} 0 \quad [\text{Law of large numbers}] \quad (4)$$

$$-P\left[\left|\frac{\sum X_i}{\sqrt{n}}\right| \in A\right] = \int_A e^{-x^2/2} dx \quad \begin{matrix} \text{Central Limit} \\ \text{Theorem} \end{matrix}$$

~ Large Deviation Principle (Cramér):

$$\log P\left(\frac{|\sum X_i|}{n} > \delta\right) \xrightarrow{n \rightarrow +\infty} -\Lambda^*(\delta)$$

1st Generalization (Giannopoulos, Gantzas):

Let μ be a measure in \mathbb{R} with compact support

$$[-x^*, x^*] \text{ even and suppose that } \lim_{x \rightarrow x^*} \frac{-\log(\mu([x, +\infty)))}{\Lambda^*(x)} = L(K)$$

If $\mu^n = \underbrace{\mu \otimes \dots \otimes \mu}_{n\text{-times}}$ (product measure), then $V \in (0, 1)$ if

$$N > e^{(1+\varepsilon) \cdot \mathbb{E}[\Lambda_{\mu^n}^*]}, \text{ then } \frac{\mathbb{E}[1_{[N, \infty]}]}{|[-x^*, x^*]^h|} \xrightarrow{n \rightarrow +\infty} 1.$$

2nd Generalization (Pafis, 2023):

The dream of Sylvester:

• Let μ be a probability measure in \mathbb{R} and suppose even.

• Let μ be a good measure. Then $V \in (0, 1)$ if $N > e^{(1+\varepsilon) \cdot \mathbb{E}[\Lambda_{\mu^n}^*]}$ then $\mathbb{E}[\mu^n(K_N)] \xrightarrow{n \rightarrow +\infty} 1$.

if $N < e^{(1-\varepsilon) \cdot \mathbb{E}[\Lambda_{\mu^n}^*]}$ then $\mathbb{E}[\mu^n(K_N)] \xrightarrow{n \rightarrow +\infty} 0$.
 (for the Gaussian measure: $\mathbb{E}[\Lambda_{\mu^n}^*] = \frac{n}{2}$)

(B., Chasapis, 2024)

*: A non trivial example for before is the case of

$$f(x) = C_P \cdot e^{-|x|^P}$$

→ - The condition (*) holds if & f there's a convex function $V: \lim_{x \rightarrow x^*} \frac{-\log(\mu([x, +\infty)))}{V(x)} = 1$.

(the tails of the measure behave log-concavely).

- If μ is any log-concave measure in \mathbb{R} and $\mu^n = \mu \otimes \mu \otimes \dots \otimes \mu$. Then, if $N > e^{(1+\varepsilon) \cdot \mathbb{E}[\lambda_{\mu^n}^*]}$ then $\mathbb{E}[\mu^n(K_N)] \rightarrow 1$ and if $N < e^{(1-\varepsilon) \cdot \mathbb{E}[\lambda_{\mu^n}^*]}$, then $\mathbb{E}[\mu^n(K_N)] \rightarrow 0$. (A measure μ in \mathbb{R}^n is log-concave if $\forall A, B$ Borel and $\lambda \in (0, 1)$)

$$\mu((1-\lambda)A + \lambda B) \geq \mu(A)^{1-\lambda} \cdot \mu(B)^\lambda$$

For all even log-concave measures: (Remark).

$\mathbb{E}[\lambda_{\mu^n}^*]$ is minimum for the measure with density $f(x) = \frac{1}{2} \cdot e^{-|x|}$

Reduction:

$$\text{If } \frac{\mathbb{V}[\lambda_{\mu}^*]}{(\mathbb{E}[\lambda_{\mu}^*])^2} \xrightarrow{n \rightarrow +\infty} 0$$

Ζητούντας μηδεμίους: [Parts III & IV].

21/7/2024

Μεσοχρηστικός Cramér, Συγκέντρωσης Βαθμών και Tukey, thresholds για τοποτυπίους υπότιτλη μέση.

Theorem: Let μ be log-concave probability measure on \mathbb{R} and $\mu_n := \underbrace{\mu \otimes \dots \otimes \mu}_{n\text{-times}}$, let X_1, X_2, \dots, X_N independent from

chosen with μ_n . If $N > e^{(1+\epsilon) \cdot \mathbb{E} A_{\mu_n}^*$, then $\mathbb{E} \mu_n(K_N) \rightarrow 1$ ($K_N := \text{conv}\{X_1, X_2, \dots, X_N\}$)

• If $N < e^{(1-\epsilon) \cdot \mathbb{E} A_{\mu_n}^*}$, then $\mathbb{E} \mu_n(K_N) \xrightarrow{n \rightarrow \infty} 0$.

where $A_{\mu_n}(\xi) = \log \mathbb{E} e^{<X, \xi>}$ and $A_{\mu_n}^*(x) = \sup_{\xi \in \mathbb{R}^n} <x, \xi> - A_{\mu_n}(\xi)$

*Reminder: μ is log-concave if $\forall A, B \in \mathcal{B}(\mathbb{R}^n)$
 $(\mu(\lambda A + (1-\lambda)B) \geq \mu(A)^{\lambda} \cdot \mu(B)^{1-\lambda}) \quad (\forall \lambda \in [0, 1]).$

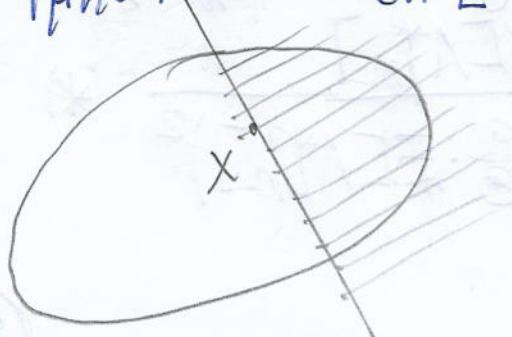
• Lemma: If ν any measure on \mathbb{R}^n and A, B Borel.

(i) $\mathbb{E} \nu(K_N) \leq \nu(A) + N \sup_{x \notin A} q_{\mu_n}(x)$

(ii) $\mathbb{E} \nu(K_N) \geq \nu(B) \cdot \left(1 - 2 \cdot \binom{N}{n} \cdot (1 - \inf_{x \in B} q_{\mu_n}(x))\right)^{N-n}$.

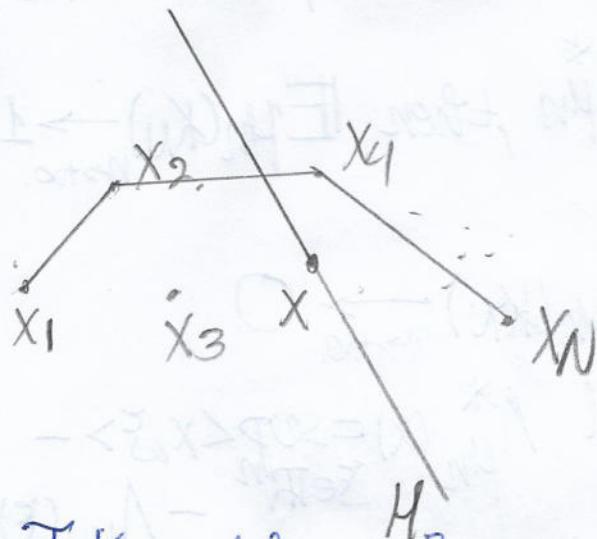
where $q_{\mu_n}(x)$ is the Tukey's halfspace depth

$q_{\mu_n}(x) := \inf \{ \text{PP} \{ <X, \xi> \geq <x, \xi> \} \}_{\xi \in H(\text{Halfspace})} = \inf_{\xi \in H(\text{Halfspace})} \text{PP} \{ X \in H \}$.



Proof of the Lemma (i):

$$\begin{aligned} \mathbb{E}[v(K_N)] &= \mathbb{E}[v(K_N \cap A)] + \mathbb{E}[v(K_N \setminus A)] \leq \\ &\leq v(A) + \mathbb{E}\left[\int_{\mathbb{R}^n \setminus A} 1_{K_N}(x) \cdot g(x) d\nu(x)\right] = v(A) + \int_{\mathbb{R}^n \setminus A} \mathbb{P}[x \in K_N] d\nu(x). \end{aligned}$$



• For any halfspace H that contains x , there is a X_i :

$X_i \in H$. So: $\{x \in K_N\} \subseteq \cup \{x_i \in H\}$

$$\Rightarrow \mathbb{P}[x \in K_N] \leq \sum \mathbb{P}[X_i \in H] = N \cdot \mathbb{P}[X_i \in H]$$

Taking the infimum over all H , then $\mathbb{P}[x \in K_N] \leq N \cdot q_{\mu_n}(x)$ and we can have (i) statement. ■

2nd Step: Choose A, B

$$A := \{x \in \mathbb{R}^n / \Lambda_{\mu_n}^*(x) \leq (1 - \frac{\varepsilon}{8}) \cdot \mathbb{E}[\Lambda_{\mu_n}^*]\}$$

$$B := \{x \in \mathbb{R}^n / \Lambda_{\mu_n}^*(x) \geq (1 + \frac{\varepsilon}{8}) \cdot \mathbb{E}[\Lambda_{\mu_n}^*]\}$$

• We have that $\max\{\mu_n(A), \mu_n(B^c)\} \leq$

$$\leq \mu_n(x \in \mathbb{R}^n / |\Lambda_{\mu_n}^*(x) - \mathbb{E}[\Lambda_{\mu_n}^*]| \geq \frac{\varepsilon}{8} \cdot \mathbb{E}[\Lambda_{\mu_n}^*])$$

≤ Chebyshov.
$$\frac{\mathbb{E}[(\Lambda_{\mu_n}^*(x) - \mathbb{E}[\Lambda_{\mu_n}^*])^2]}{(\frac{\varepsilon}{8} \cdot \mathbb{E}[\Lambda_{\mu_n}^*])^2} = \frac{\mathbb{V}[\Lambda_{\mu_n}^*]}{(\frac{\varepsilon}{8} \cdot \mathbb{E}[\Lambda_{\mu_n}^*])^2} \quad (*)$$

[Fact: $\Lambda_{\mu_n}^*(x_1, \dots, x_n) = \Lambda_\mu^*(x_1) + \dots + \Lambda_\mu^*(x_n)$]

Therefore, $(*) = \frac{n \cdot \mathbb{V}[\Lambda_\mu^*]}{\left(\frac{\epsilon}{8} \cdot n \mathbb{E}[\Lambda_\mu^*]\right)^2} = \frac{1}{n \cdot \frac{\epsilon^2}{8^2}} \cdot \frac{\mathbb{V}[\Lambda_\mu^*]}{(n \mathbb{E}[\Lambda_\mu^*])^2}$

• Relation between q_{μ_n} and $\Lambda_{\mu_n}^*$ Chernoff.

$$\text{Fix } \xi \in \mathbb{R}^n \quad \mathbb{P}[E[X, \xi] \geq E[X, \xi]] \leq \mathbb{E}[e^{E[X, \xi] - E[X, \xi]}]$$

$$\text{Taking inf } q_{\mu_n}(x) \leq e^{-\Lambda_{\mu_n}^*(x)} \Rightarrow$$

$$\Rightarrow \Lambda_{\mu_n}^*(x) \leq \ln \frac{1}{q_{\mu_n}(x)}$$

► Thm: If μ is probability measure on \mathbb{R} which is log-concave (continuous or discrete). Then $\forall \epsilon \in (0, 1)$ and $x \in \overline{\text{supp}(\mu)}$, $\Lambda_\mu^*(x) \geq (1-\epsilon) \cdot \ln \frac{1}{q_{\mu_n}(x)} + \log \frac{\epsilon}{2^{1-\epsilon}}$

The inequality is sharp.

Proof: • q_μ is log-concave:

$$\mathbb{P}[E[X, \xi] \geq (1-\epsilon)x + \epsilon y, \xi] \geq \mathbb{P}[E[X, \xi] \geq E[X, \xi]]^{1-\epsilon}$$

• $\mathbb{P}[E[X, \xi] \geq E[X, \xi]] \geq q(x)^{1-\epsilon} \cdot q(y)^\epsilon$. Taking the inf we have the log-concavity.

• Fix $x \in \text{supp}(\mu)$ and $\epsilon \in (0, 1)$. Consider $g(y) = (1-\epsilon) \ln \frac{1}{q_\mu(y)}$ (which is convex).

(*) Discrete log-concavity: $(p_{CK})^2 \geq p_{(K+1)} \cdot p_{(K-1)}$.

Therefore, $g(x) = l(x)$ and.

$g(y) \geq l(y) \quad \forall y \neq x$.

$$\Rightarrow \mathbb{E}[e^{l(X)}] \leq \mathbb{E}[e^{g(X)}],$$

$$l(x) - \log \mathbb{E}[e^{l(X)}] \geq g(x) - \log \mathbb{E}[e^{g(X)}] \quad l(y) = y \cdot t + \theta.$$

$$\Rightarrow xt + \theta - \log(\mathbb{E}[e^{tx+\theta}]) \geq g(x) - \log \mathbb{E}[e^{g(x)}]$$

$$\Rightarrow xt - \Lambda(t) \geq (1-\varepsilon) \cdot \ln \frac{1}{q_\mu(x)} - \log \mathbb{E}\left[\left(\frac{1}{q_\mu(x)}\right)^{1-\varepsilon}\right]$$

$$\Rightarrow \Lambda_\mu^*(x) \geq \dots$$

Fact: If $F(x) = \mathbb{P}[X \leq x]$ and F is continuous then the r.v. $F(X)$ is uniformly distributed in $[0, 1]$.

$$\mathbb{P}[F(X) \leq x] = \mathbb{P}[X \leq F^{-1}(x)] = x.$$

Moreover $q(x) = \min\{F(x), 1-F(x)\}$, therefore

$$q(x) = \min\{U, 1-U\}, \quad U \sim \text{Unif}[0, 1].$$

$$\mathbb{E}[q(x)^p] = \mathbb{E}[(\min\{U, 1-U\})^p] = \int_0^{1/2} u^p du + \int_{1/2}^1 (1-u)^p du$$

$$\Rightarrow \mathbb{E}[q(x)^p] = \frac{1}{2p(1+p)}$$

$$\Lambda_\mu^*(x) \geq (1-\varepsilon) \cdot \log \frac{1}{q_\mu(x)} + \log \frac{\varepsilon}{2^{1-\varepsilon}}, \quad \varepsilon \in (0, 1).$$

Let $A := \{x \in \mathbb{R}^n / q_\mu(x) < \frac{1}{2e}\}$, then choose, for $x \in A$,

$$\varepsilon_e = \left(\log \frac{1}{2q(x)}\right)^{-1} \in (0, 1).$$

$$\forall x \in A: \Lambda_\mu^*(x) \geq \log\left(\frac{1}{2q(x) \cdot \log\left(\frac{1}{2q(x)}\right)}\right) - 1.$$

$$\Rightarrow \mathbb{E}[\Lambda_{\mu}^*] \geq \mathbb{E}_{\lambda}[\Lambda_{\mu}^*] \geq \mathbb{E}\left[\log \frac{1}{2q(x) \cdot \log\left(\frac{1}{2q(x)}\right)} - 1\right] \sim 0.14.$$

Upper Bound: $\mathbb{E}[\Lambda_{\mu}^{*2}] \leq \mathbb{E}\left[\log\left(\frac{1}{q(x)}\right)\right]^2 \sim 3.8.$

Thm: For all continuous probability measures: the one dimensional marginals are log-concave, we have that $\forall x$ and $\epsilon \in (0, 1)$.

$$\Lambda_{\mu}^*(x) \geq (1-\epsilon) \cdot \log \frac{1}{q_{\mu}(x)} + \log \frac{\epsilon}{2^{1-\epsilon}}.$$

* We can't prove this theorem in the discrete case.

$$\Lambda_{\mu}^*(x) = \sup_{\xi \in \mathbb{R}^n} (\langle x, \xi \rangle - \Lambda_{\mu}(\xi)) = \sup_{(t, \xi) \in \mathbb{R} \times \mathbb{S}^{n-1}} (t \cdot \langle x, \xi \rangle -$$

$$- \Lambda(t \cdot \langle x, \xi \rangle) = \sup_{\xi \in \mathbb{S}^{n-1}} \sup_t (t \cdot \langle x, \xi \rangle - \Lambda_{\langle x, \xi \rangle}(t)) =$$

$$= \sup_{\xi \in \mathbb{S}^{n-1}} \Lambda_{\langle x, \xi \rangle}^*(\langle x, \xi \rangle).$$

For fixed $\xi \in \mathbb{S}^{n-1}$, $\Lambda_{\langle x, \xi \rangle}^*(\langle x, \xi \rangle) \geq$

$$(1-\epsilon) \cdot \log \frac{1}{q_{\langle x, \xi \rangle}(\langle x, \xi \rangle)} + \log \frac{\epsilon}{2^{1-\epsilon}}.$$

$$\begin{aligned}
 N \cdot q_{\mu_n}(x) &\leq N \cdot e^{-\lambda_{\mu_n}^*(x)} \leq N \cdot e^{-(1-\frac{\varepsilon}{8}) \cdot \mathbb{E}[\lambda_{\mu_n}^*]} \leq \\
 &\leq e^{(1-\varepsilon) \cdot \mathbb{E}[\lambda_{\mu_n}^*] - (1-\frac{\varepsilon}{8}) \cdot \mathbb{E}[\lambda_{\mu_n}^*]} = e^{-\frac{7\varepsilon}{8} \mathbb{E}[\lambda_{\mu_n}^*]} \\
 &= e^{-\frac{7\varepsilon}{8} \cdot n \mathbb{E}[\lambda_{\mu_n}^*]}.
 \end{aligned}$$

The End!!!