

Asymptotic shape of a random polytope in a convex body

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Abstract

Let K be an isotropic convex body in \mathbb{R}^n and let $Z_q(K)$ be the L_q -centroid body of K . For every $N > n$ consider the random polytope $K_N := \text{conv}\{x_1, \dots, x_N\}$ where x_1, \dots, x_N are independent random points, uniformly distributed in K . We prove that a random K_N is “asymptotically equivalent” to $Z_{\lfloor \ln(N/n) \rfloor}(K)$ in the following sense: there exist absolute constants $\rho_1, \rho_2 > 0$ such that, for all $\beta \in (0, \frac{1}{2}]$ and all $N \geq N(n, \beta)$, one has:

(i) $K_N \supseteq c(\beta)Z_q(K)$ for every $q \leq \rho_1 \ln(N/n)$, with probability greater than $1 - c_1 \exp(-c_2 N^{1-\beta} n^\beta)$.

(ii) For every $q \geq \rho_2 \ln(N/n)$, the expected mean width $\mathbb{E}[w(K_N)]$ of K_N is bounded by $c_3 w(Z_q(K))$.

As an application we show that the volume radius $|K_N|^{1/n}$ of a random K_N satisfies the bounds $c_4 \frac{\sqrt{\ln(2N/n)}}{\sqrt{n}} \leq |K_N|^{1/n} \leq c_5 L_K \frac{\sqrt{\ln(2N/n)}}{\sqrt{n}}$ for all $N \leq \exp(n)$.

1 Introduction

Let K be a convex body of volume 1 in \mathbb{R}^n . For every $q \geq 1$ we define the L_q -centroid body $Z_q(K)$ of K by its support function:

$$(1.1) \quad h_{Z_q(K)}(x) = \|\langle \cdot, x \rangle\|_q := \left(\int_K |\langle y, x \rangle|^q dy \right)^{1/q}.$$

The aim of this article is to provide some precise quantitative information on the “asymptotic shape” of a random polytope $K_N = \text{conv}\{x_1, \dots, x_N\}$ spanned by N independent random points x_1, \dots, x_N uniformly distributed in K . Our approach is to compare K_N with the L_q -centroid body $Z_q(K)$ of K for $q \simeq \ln(N/n)$.

The origin of our work is in the study of the behavior of symmetric random ± 1 -polytopes, the absolute convex hulls of random subsets of the discrete cube $E_2^n = \{-1, 1\}^n$. The natural way to produce these random polytopes is to fix $N > n$ and to consider the convex hull $K_{n,N} = \text{conv}\{\pm \vec{X}_1, \dots, \pm \vec{X}_N\}$ of N independent random points $\vec{X}_1, \dots, \vec{X}_N$, uniformly distributed over E_2^n . It turns out

(see [9]) that a random polytope $K_{n,N}$ has the largest possible volume among all ± 1 -polytopes with N vertices, at every scale of n and N . This is a consequence of the following fact: If $n \geq n_0$ and if $N \geq n(\ln n)^2$, then

$$(1.2) \quad K_{n,N} \supseteq c \left(\sqrt{\ln(N/n)} B_2^n \cap B_\infty^n \right)$$

with probability greater than $1 - e^{-n}$, where $c > 0$ is an absolute constant, B_2^n is the Euclidean unit ball in \mathbb{R}^n and $B_\infty^n = [-1, 1]^n$.

In [16], Litvak, Pajor, Rudelson, and Tomczak-Jaegermann worked in a more general setting which contains the previous Bernoulli model and the Gaussian model; let $K_{n,N}$ be the absolute convex hull of the rows of the random matrix $\Gamma_{n,N} = (\xi_{ij})_{1 \leq i \leq N, 1 \leq j \leq n}$, where ξ_{ij} are independent symmetric random variables satisfying certain conditions ($\|\xi_{ij}\|_{L^2} \geq 1$ and $\|\xi_{ij}\|_{L^{\psi_2}} \leq \rho$ for some $\rho \geq 1$, where $\|\cdot\|_{L^{\psi_2}}$ is the Orlicz norm corresponding to the function $\psi_2(t) = e^{t^2} - 1$). For this larger class of random polytopes, the estimates from [9] were generalized and improved in two ways: the paper [16] provides estimates for all $N \geq (1+\delta)n$, where $\delta > 0$ can be as small as $1/\ln n$, and establishes the following inclusion: for every $0 < \beta < 1$,

$$(1.3) \quad K_{n,N} \supseteq c(\rho) \left(\sqrt{\beta \ln(N/n)} B_2^n \cap B_\infty^n \right)$$

with probability greater than $1 - \exp(-c_1 n^\beta N^{1-\beta}) - \exp(-c_2 N)$. The proof in [16] is based on a lower bound of the order of \sqrt{N} for the smallest singular value of the random matrix $\Gamma_{n,N}$ with probability greater than $1 - \exp(-cN)$.

In a sense, both works correspond to the study of the size of a random polytope $K_N = \text{conv}\{x_1, \dots, x_N\}$ spanned by N independent random points x_1, \dots, x_N uniformly distributed in the unit cube $Q_n := [-1/2, 1/2]^n$. The connection of the estimates (1.2) and (1.3) with L_q -centroid bodies comes from the following observation.

Remark. For $x \in \mathbb{R}^n$ and $t > 0$, define

$$(1.4) \quad K_{1,2}(x, t) := \inf \{ \|u\|_1 + t \|x - u\|_2 : u \in \mathbb{R}^n \}.$$

If we write $(x_j^*)_{j \leq n}$ for the decreasing rearrangement of $(|x_j|)_{j \leq n}$ we have Holmstedt's approximation formula

$$(1.5) \quad \frac{1}{c} K_{1,2}(x, t) \leq \sum_{j=1}^{\lfloor t^2 \rfloor} x_j^* + t \left(\sum_{j=\lfloor t^2 \rfloor + 1}^n (x_j^*)^2 \right)^{1/2} \leq K_{1,2}(x, t)$$

where $c > 0$ is an absolute constant (see [14]). Now, for any $\alpha \geq 1$ define $C(\alpha) = \alpha B_2^n \cap B_\infty^n$. Then,

$$(1.6) \quad h_{C(\alpha)}(\theta) = K_{1,2}(\theta, \alpha)$$

for every $\theta \in S^{n-1}$. On the other hand,

$$(1.7) \quad \|\langle \cdot, \theta \rangle\|_{L^q(Q_n)} \simeq \sum_{j \leq q} \theta_j^* + \sqrt{q} \left(\sum_{q < j \leq n} (\theta_j^*)^2 \right)^{1/2}$$

for every $q \geq 1$ (see, for example, [6]). In other words,

$$(1.8) \quad C(\sqrt{q}) \simeq Z_q(Q_n)$$

where $Z_q(K)$ is the L_q -centroid body of K . This shows that (1.3) or (1.2) can be written in the form

$$(1.9) \quad K_{n,N} \supseteq c(\rho) Z_{\beta \ln(N/n)}(Q_n).$$

This observation leads us to consider a random polytope $K_N = \text{conv}\{x_1, \dots, x_N\}$ spanned by N independent random points x_1, \dots, x_N uniformly distributed in an isotropic convex body K and try to compare K_N with $Z_q(K)$ for a suitable value $q = q(N, n) \simeq \ln(N/n)$. Our first main result states that an analogue of (1.9) holds true in full generality.

Theorem 1.1 *Let $\beta \in (0, 1/2]$ and $\gamma > 1$. If*

$$(1.10) \quad N \geq N(\gamma, n) = c\gamma n,$$

where $c > 0$ is an absolute constant, for every isotropic convex body K in \mathbb{R}^n we have

$$(1.11) \quad K_N \supseteq c_1 Z_q(K) \text{ for all } q \leq c_2 \beta \ln(N/n),$$

with probability greater than

$$(1.12) \quad 1 - \exp(-c_3 N^{1-\beta} n^\beta) - \mathbb{P}(\|\Gamma : \ell_2^n \rightarrow \ell_2^N\| \geq \gamma L_K \sqrt{N}),$$

where $\Gamma : \ell_2^n \rightarrow \ell_2^N$ is the random operator $\Gamma(y) = (\langle x_1, y \rangle, \dots, \langle x_N, y \rangle)$ defined by the vertices x_1, \dots, x_N of K_N .

The proof of Theorem 1.1 is given in Section 2, where we also collect what is known about the probability $\mathbb{P}(\|\Gamma : \ell_2^n \rightarrow \ell_2^N\| \geq \gamma L_K \sqrt{N})$ which appears in (1.12).

It should be emphasized that a reverse inclusion of the form $K_N \subseteq c_4 Z_q(K)$ cannot be expected with probability close to 1, unless q is of the order of n . This follows by a simple volume argument which makes use of the upper estimate of Paouris (see [20]) for the volume of $Z_q(K)$ and is presented in Section 3. However, one can easily see that K_N is “weakly sandwiched” between $Z_{q_i}(K)$ ($i = 1, 2$), where $q_i \simeq \ln(N/n)$, in the following sense:

Proposition 1.2 *For every $\alpha > 1$ one has*

$$(1.13) \quad \mathbb{E} [\sigma(\theta : (h_{K_N}(\theta) \geq \alpha h_{Z_q(K)}(\theta)))] \leq N\alpha^{-q}.$$

This shows that if $q \geq c_5 \ln(N/n)$ then, for most $\theta \in S^{n-1}$, one has $h_{K_N}(\theta) \leq c_6 h_{Z_q(K)}(\theta)$. It follows that several geometric parameters of K_N , e.g. the mean width, are controlled by the corresponding parameter of $Z_{[\ln(N/n)]}(K)$.

As an application, we discuss the volume radius of K_N : Let K be a convex body of volume 1 in \mathbb{R}^n . The question to estimate the expected volume radius

$$(1.14) \quad \mathbb{E}(K, N) = \int_K \cdots \int_K |\text{conv}(x_1, \dots, x_N)|^{1/n} dx_N \cdots dx_1$$

of K_N was studied in [12] where it was proved that for every isotropic convex body K in \mathbb{R}^n and every $N \geq n + 1$,

$$(1.15) \quad \mathbb{E}(B(n), N) \leq \mathbb{E}(K, N) \leq cL_K \frac{\ln(2N/n)}{\sqrt{n}},$$

where $B(n)$ is a ball of volume 1. This estimate is rather weak for large values of N : a strong conjecture is that

$$(1.16) \quad \mathbb{E}(K, N) \simeq \min \left\{ \frac{\sqrt{\ln(2N/n)}}{\sqrt{n}}, 1 \right\} L_K$$

for every $N \geq n + 1$. This was verified in [10] in the unconditional case, where it was also shown that the general problem is related to the “ ψ_2 -behavior” of linear functionals on isotropic convex bodies. Using a recent result of G. Paouris [21] on the negative moments of the support function of $h_{Z_q(K)}$ we can settle the question for the full range of values of N .

Theorem 1.3 *For every $N \leq \exp(n)$, one has*

$$(1.17) \quad c_4 \frac{\sqrt{\ln(2N/n)}}{\sqrt{n}} \leq |K_N|^{1/n} \leq c_5 L_K \frac{\sqrt{\ln(2N/n)}}{\sqrt{n}}$$

with probability greater than $1 - \frac{1}{N}$, where $c_4, c_5 > 0$ are absolute constants.

Notation and terminology. We work in \mathbb{R}^n , which is equipped with a Euclidean structure $\langle \cdot, \cdot \rangle$. We denote by $\|\cdot\|_2$ the corresponding Euclidean norm, and write B_2^n for the Euclidean unit ball, and S^{n-1} for the unit sphere. Volume is denoted by $|\cdot|$. We write ω_n for the volume of B_2^n and σ for the rotationally invariant probability measure on S^{n-1} . We also write \bar{A} for the homothetic image of volume 1 of a convex body $A \subseteq \mathbb{R}^n$, i.e. $\bar{A} := \frac{A}{|A|^{1/n}}$.

A convex body is a compact convex subset C of \mathbb{R}^n with non-empty interior. We say that C is symmetric if $-x \in C$ whenever $x \in C$. We say that C has center of mass at the origin if $\int_C \langle x, \theta \rangle dx = 0$ for every $\theta \in S^{n-1}$. The support function $h_C : \mathbb{R}^n \rightarrow \mathbb{R}$ of C is defined by $h_C(x) = \max\{\langle x, y \rangle : y \in C\}$. The mean width of C is defined by

$$(1.18) \quad w(C) = \int_{S^{n-1}} h_C(\theta) \sigma(d\theta).$$

The radius of C is the quantity $R(C) = \max\{\|x\|_2 : x \in C\}$, and the polar body C° of C is

$$(1.19) \quad C^\circ := \{y \in \mathbb{R}^n : \langle x, y \rangle \leq 1 \text{ for all } x \in C\}.$$

Whenever we write $a \simeq b$, we mean that there exist universal constants $c_1, c_2 > 0$ such that $c_1 a \leq b \leq c_2 a$. The letters $c, c', c_1, c_2 > 0$ etc., denote universal positive constants which may change from line to line.

A convex body K in \mathbb{R}^n is called isotropic if it has volume $|K| = 1$, center of mass at the origin, and there is a constant $L_K > 0$ such that

$$(1.20) \quad \int_K \langle x, \theta \rangle^2 dx = L_K^2$$

for every θ in the Euclidean unit sphere S_2^{n-1} . For every convex body K in \mathbb{R}^n there exists an affine transformation T of \mathbb{R}^n such that $T(K)$ is isotropic. Moreover, if we ignore orthogonal transformations, this isotropic image is unique, and hence, the isotropic constant L_K is an invariant of the affine class of K . We refer to [18] and [8] for more information on isotropic convex bodies.

2 The main inclusion

In this Section we prove Theorem 1.1. Let K be an isotropic convex body in \mathbb{R}^n . For every $q \geq 1$ consider the L_q -centroid body $Z_q(K)$ of K ; recall that

$$(2.1) \quad h_{Z_q(K)}(x) = \|\langle \cdot, x \rangle\|_q := \left(\int_K |\langle y, x \rangle|^q dy \right)^{1/q}.$$

Since $|K| = 1$, we readily see that $Z_1(K) \subseteq Z_p(K) \subseteq Z_q(K) \subseteq Z_\infty(K)$ for every $1 \leq p \leq q \leq \infty$, where $Z_\infty(K) = \text{conv}\{K, -K\}$. On the other hand, one has the reverse inclusions

$$(2.2) \quad Z_q(K) \subseteq \frac{cq}{p} Z_p(K)$$

for every $1 \leq p < q < \infty$, as a consequence of the ψ_1 -behavior of $y \mapsto \langle y, x \rangle$. Observe that $Z_q(K)$ is always symmetric, and $Z_q(TK) = T(Z_q(K))$ for every $T \in SL(n)$ and $q \in [1, \infty]$. Also, if K has its center of mass at the origin, then $Z_q(K) \supseteq cZ_\infty(K)$ for all $q \geq n$, where $c > 0$ is an absolute constant. We refer to [8] for proofs of these assertions and further information.

Lemma 2.1 *Let $0 < t < 1$. For every $\theta \in S^{n-1}$ one has*

$$(2.3) \quad \mathbb{P}(\{x \in K : |\langle x, \theta \rangle| \geq t \|\langle \cdot, \theta \rangle\|_q\}) \geq \frac{(1-t^q)^2}{C^q}.$$

Proof. We apply the Paley-Zygmund inequality

$$(2.4) \quad \mathbb{P}(g(x) \geq t^q \mathbb{E}(g)) \geq (1 - t^q)^2 \frac{[\mathbb{E}(g)]^2}{\mathbb{E}(g^2)}$$

for the function $g(x) = |\langle x, \theta \rangle|^q$. Since, by (2.2),

$$(2.5) \quad \mathbb{E}(g^2) = \mathbb{E}|\langle x, \theta \rangle|^{2q} \leq C^q (\mathbb{E}|\langle x, \theta \rangle|^q)^2 = C^q [\mathbb{E}(g)]^2$$

for some absolute constant $C > 0$, the lemma is proved. \square

Lemma 2.2 *For every $\sigma \subseteq \{1, \dots, N\}$ and any $\theta \in S^{n-1}$ one has*

$$(2.6) \quad \mathbb{P}\left(\{\vec{X} = (x_1, \dots, x_N) \in K^N : \max_{j \in \sigma} |\langle x_j, \theta \rangle| \leq \frac{1}{2} \|\langle \cdot, \theta \rangle\|_q\}\right) \leq \exp(-|\sigma|/(4C^q)),$$

where $C > 0$ is an absolute constant.

Proof. Applying Lemma 2.1 with $t = 1/2$ we see that

$$\begin{aligned} \mathbb{P}\left(\max_{j \in \sigma} |\langle x_j, \theta \rangle| \leq \frac{1}{2} \|\langle \cdot, \theta \rangle\|_q\right) &= \prod_{j \in \sigma} \mathbb{P}\left(|\langle x_j, \theta \rangle| \leq \frac{1}{2} \|\langle \cdot, \theta \rangle\|_q\right) \\ &\leq \left(1 - \frac{1}{4C^q}\right)^{|\sigma|} \\ &\leq \exp(-|\sigma|/(4C^q)), \end{aligned}$$

since $1 - v < e^{-v}$ for every $v > 0$. \square

Proof of Theorem 1.1. Let $\Gamma : \ell_2^n \rightarrow \ell_2^N$ be the random operator defined by

$$(2.7) \quad \Gamma(y) = (\langle x_1, y \rangle, \dots, \langle x_N, y \rangle).$$

We modify an idea from [16]. Define $m = \lceil 8(N/n)^{2\beta} \rceil$ and $k = \lceil N/m \rceil$. Fix a partition $\sigma_1, \dots, \sigma_k$ of $\{1, \dots, N\}$ with $m \leq |\sigma_i|$ for all $i = 1, \dots, k$ and define the norm

$$(2.8) \quad \|u\|_0 = \frac{1}{k} \sum_{i=1}^k \|P_{\sigma_i}(u)\|_\infty.$$

Since

$$(2.9) \quad h_{K_N}(z) = \max_{1 \leq j \leq N} |\langle x_j, z \rangle| \geq \|P_{\sigma_i} \Gamma(z)\|_\infty$$

for all $z \in \mathbb{R}^n$ and $i = 1, \dots, k$, we observe that

$$(2.10) \quad h_{K_N}(z) \geq \|\Gamma(z)\|_0.$$

If $z \in \mathbb{R}^n$ and $\|\Gamma(z)\|_0 < \frac{1}{4}\|\langle \cdot, z \rangle\|_q$, then, Markov's inequality implies that there exists $I \subset \{1, \dots, k\}$ with $|I| > k/2$ such that $\|P_{\sigma_i}\Gamma(z)\|_\infty < \frac{1}{2}\|\langle \cdot, z \rangle\|_q$, for all $i \in I$. It follows that, for fixed $z \in S^{n-1}$ and $\alpha \geq 1$,

$$\begin{aligned}
& \mathbb{P}\left(\|\Gamma(z)\|_0 < \frac{1}{4}\|\langle \cdot, z \rangle\|_q\right) \\
& \leq \sum_{|I|=\lceil(k+1)/2\rceil} \mathbb{P}\left(\|P_{\sigma_i}\Gamma(z)\|_\infty < \frac{1}{2}\|\langle \cdot, z \rangle\|_q, \text{ for all } i \in I\right) \\
& \leq \sum_{|I|=\lceil(k+1)/2\rceil} \prod_{i \in I} \mathbb{P}\left(\|P_{\sigma_i}\Gamma(z)\|_\infty < \frac{1}{2}\|\langle \cdot, z \rangle\|_q\right) \\
& \leq \sum_{|I|=\lceil(k+1)/2\rceil} \prod_{i \in I} \exp(-|\sigma_i|/(4C^q)) \\
& \leq \binom{k}{\lceil(k+1)/2\rceil} \exp(-c_1 km/C^q) \\
& \leq \exp(k \ln 2 - c_1 km/C^q).
\end{aligned}$$

Choosing

$$(2.11) \quad q \simeq \beta \ln(N/n)$$

we see that

$$(2.12) \quad \mathbb{P}\left(\|\Gamma(z)\|_0 < \frac{1}{4}\|\langle \cdot, z \rangle\|_q\right) \leq \exp(-c_2 N^{1-\beta} n^\beta).$$

Let $S = \{z : \|\langle \cdot, z \rangle\|_q/2 = 1\}$ and consider a δ -net U of S with cardinality $|U| \leq (3/\delta)^n$. For every $u \in U$ we have

$$(2.13) \quad \mathbb{P}\left(\|\Gamma(u)\|_0 < \frac{1}{2}\right) \leq \exp(-c_2 N^{1-\beta} n^\beta),$$

and hence,

$$(2.14) \quad \mathbb{P}\left(\bigcup_{u \in U} \left\{\|\Gamma(u)\|_0 < \frac{1}{2}\right\}\right) \leq \exp(n \ln(3/\delta) - c_2 N^{1-\beta} n^\beta).$$

Fix $\gamma > 1$ and set

$$(2.15) \quad \Omega_\gamma = \{\Gamma : \|\Gamma : \ell_2^n \rightarrow \ell_2^N\| \leq \gamma L_K \sqrt{N}\}.$$

Since $Z_q(K) \supseteq cL_K B_2^n$, we have

$$(2.16) \quad \|\Gamma(z)\|_0 \leq \frac{1}{\sqrt{k}} \|\Gamma(z)\|_2 \leq c\gamma L_K \sqrt{N/k} \|z\|_2 \leq c\gamma \sqrt{N/k} \|\langle \cdot, z \rangle\|_q$$

for all $z \in \mathbb{R}^n$ and all Γ in Ω_γ .

Let $z \in S$. There exists $u \in U$ such that $\frac{1}{2}\|\langle \cdot, z - u \rangle\|_q < \delta$, which implies that

$$(2.17) \quad \|\Gamma(u)\|_0 \leq \|\Gamma(z)\|_0 + c\gamma\delta\sqrt{N/k}$$

on Ω_γ . Now, choose $\delta = \sqrt{k/N}/(4c\gamma)$. Then,

$$\begin{aligned} & \mathbb{P}(\{\Gamma \in \Omega_\gamma : \exists z \in \mathbb{R}^n : \|\Gamma(z)\|_0 \leq \|\langle \cdot, z \rangle\|_q/8\}) \\ &= \mathbb{P}(\{\Gamma \in \Omega_\gamma : \exists z \in S : \|\Gamma(z)\|_0 \leq 1/4\}) \\ &\leq \mathbb{P}(\{\Gamma \in \Omega_\gamma : \exists u \in U : \|\Gamma(u)\|_0 \leq 1/2\}) \\ &\leq \exp\left(n \ln(12c\gamma\sqrt{N/k}) - c_2N^{1-\beta}n^\beta\right) \\ &\leq \exp(-c_3N^{1-\beta}n^\beta) \end{aligned}$$

provided that N is large enough. Since $h_{K_N}(z) \geq \|\Gamma(z)\|_0$ for every $z \in \mathbb{R}^n$, we get that $K_N \supseteq cZ_q(K)$ with probability greater than $1 - \exp(-c_4N^{1-\beta}n^\beta) - \mathbb{P}(\|\Gamma : \ell_2^n \rightarrow \ell_2^N\| \geq \gamma L_K \sqrt{N})$.

We now analyze the restriction for N ; we need $n \ln(12c_4\gamma\sqrt{N/k}) \leq CN^{1-\beta}n^\beta$ for some suitable constant $C > 0$. Assuming

$$(2.18) \quad N \geq 12c\gamma n,$$

and since $\beta \in (0, \frac{1}{2}]$, using the definitions of k and m we see that it is enough to guarantee

$$\ln(N/n) \leq C\sqrt{N/n},$$

which is valid if $N/n \geq c_5$ for a suitable absolute constant $c_5 > 0$. We get the result taking (2.18) into account. \square

Remark 2.3 The statement of Theorem 1.1 raises the question to estimate the probability

$$(2.19) \quad \mathbb{P}(\Omega_\gamma) = \mathbb{P}(\|\Gamma : \ell_2^n \rightarrow \ell_2^N\| \geq \gamma L_K \sqrt{N}).$$

In [16] it was proved that if $T_{n,N} = (\xi_{ij})_{1 \leq i \leq N, 1 \leq j \leq n}$ is a random matrix, where ξ_{ij} are independent symmetric random variables satisfying $\|\xi_{ij}\|_{L^2} \geq 1$ and $\|\xi_{ij}\|_{L^{\psi_2}} \leq \rho$ for some $\rho \geq 1$, then $\mathbb{P}(\Omega_\gamma) \leq \exp(-c(\rho, \gamma)N)$. In our case, Γ is a random $N \times n$ matrix whose rows are N uniform random points from an isotropic convex body K in \mathbb{R}^n . Then, the question is to estimate the probability that, N random points x_1, \dots, x_N from K satisfy

$$(2.20) \quad \frac{1}{N} \sum_{j=1}^N \langle x_j, \theta \rangle^2 \leq \gamma^2 L_K^2$$

for all $\theta \in S^{n-1}$. This is related to the following well-studied question of Kannan, Lovász and Simonovits [15] which has its origin in the problem of finding a fast

algorithm for the computation of the volume of a given convex body: given $\delta, \varepsilon \in (0, 1)$, find the smallest positive integer $N_0(n, \delta, \varepsilon)$ so that if $N \geq N_0$ then with probability greater than $1 - \delta$ one has

$$(2.21) \quad (1 - \varepsilon)L_K^2 \leq \frac{1}{N} \sum_{j=1}^N \langle x_j, \theta \rangle^2 \leq (1 + \varepsilon)L_K^2$$

for all $\theta \in S^{n-1}$. In [15] it was proved that one can have $N_0 \simeq c(\delta, \varepsilon)n^2$, which was later improved to $N_0 \simeq c(\delta, \varepsilon)n(\ln n)^3$ by Bourgain [2] and to $N_0 \simeq c(\delta, \varepsilon)n(\ln n)^2$ by Rudelson [24]. One can actually check (see [11]) that this last estimate can be obtained by Bourgain's argument if we also use Alesker's concentration inequality. For subsequent developments, see see, for instance, [20], [13], [17] and [1].

Here, we are only interested in the upper bound of (2.21); actually, we need an isomorphic version of this upper estimate, since we are allowed to choose an absolute constant $\gamma \gg 1$ in (2.20). An application of the main result of [17] to the isotropic case gives such an estimate: If $N \geq c_1 n \ln^2 n$, then

$$(2.22) \quad \mathbb{P}(\|\Gamma : \ell_2^n \rightarrow \ell_2^N\| \geq \gamma L_K \sqrt{N}) \leq \exp\left(-c_2 \gamma \left(\frac{N}{(\ln N)(n \ln n)}\right)^{1/4}\right).$$

A slightly better estimate can be extracted from the work of Guédon and Rudelson in [13]. It should be emphasized that this term does not allow us to fully exploit the second term $\exp(-c_3 N^{1-\beta} n^\beta)$ in the probability estimate of Theorem 1.1. However, it is not clear if it is optimal.

Remark 2.4 G. Paouris and E. Werner [22] have recently studied the relation between the family of L_q -centroid bodies and the family of floating bodies of a convex body K . Given $\delta \in (0, \frac{1}{2}]$, the floating body K_δ of K is the intersection of all halfspaces whose defining hyperplanes cut off a set of volume δ from K . It was observed in [18] that K_δ is isomorphic to an ellipsoid as long as δ stays away from 0. In [22] it is proved that

$$(2.23) \quad c_1 Z_{\ln(1/\delta)}(K) \subseteq K_\delta \subseteq c_2 Z_{\ln(1/\delta)}(K)$$

where $c_1, c_2 > 0$ are absolute constants. From Theorem 1.1 it follows that if K is isotropic and if, for example, $N \geq n^2$ then

$$(2.24) \quad K_N \supseteq c_3 K_{1/N}$$

with probability greater than $1 - o_n(1)$, where $c_3 > 0$ is an absolute constant. This fact should be compared with the following well-known result from [3]: for any convex body K in \mathbb{R}^n one has $c|K_{1/N}| \leq \mathbb{E}|K_N| \leq c_n|K_{1/N}|$ (where the constant on the left is absolute and the right hand side inequality holds true with a constant c_n depending on the dimension, for N large enough; the critical value of N is exponential in n).

2.1 Unconditional case

In this subsection we consider separately the case of unconditional convex bodies: we assume that K is centrally symmetric and that, after a linear transformation, the standard orthonormal basis $\{e_1, \dots, e_n\}$ of \mathbb{R}^n is a 1-unconditional basis for $\|\cdot\|_K$, i.e. for every choice of real numbers t_1, \dots, t_n and every choice of signs $\varepsilon_j = \pm 1$,

$$(2.25) \quad \|\varepsilon_1 t_1 e_1 + \dots + \varepsilon_n t_n e_n\|_K = \|t_1 e_1 + \dots + t_n e_n\|_K.$$

Then, a diagonal operator brings K to the isotropic position. It is also known that the isotropic constant of an unconditional convex body K satisfies $L_K \simeq 1$.

Bobkov and Nazarov have proved that $K \supseteq c_2 Q_n$, where $Q_n = [-\frac{1}{2}, \frac{1}{2}]^n$ (see [4]). The following argument of R. Latała (private communication) shows that the family of L_q -centroid bodies of the cube Q_n is extremal in the sense that $Z_q(K) \supseteq c Z_q(Q_n)$ for all $q \geq 1$, where $c > 0$ is an absolute constant: Let $\varepsilon_1, \varepsilon_2, \dots, \varepsilon_n$ be independent and identically distributed ± 1 random variables, defined on some probability space $(\Omega, \mathcal{F}, \mathbb{P})$, with distribution $\mathbb{P}(\varepsilon_i = 1) = \mathbb{P}(\varepsilon_i = -1) = \frac{1}{2}$. For every $\theta \in S^{n-1}$, by the unconditionality of K , Jensen's inequality and the contraction principle, one has

$$\begin{aligned} \|\langle \cdot, \theta \rangle\|_{L^q(K)} &= \left(\int_K \left| \sum_{i=1}^n \theta_i x_i \right|^q dx \right)^{1/q} = \left(\int_{\Omega} \int_K \left| \sum_{i=1}^n \theta_i \varepsilon_i |x_i| \right|^q dx d\mathbb{P}(\varepsilon) \right)^{1/q} \\ &\geq \left(\int_{\Omega} \left| \sum_{i=1}^n \theta_i \varepsilon_i \int_K |x_i| dx \right|^q d\mathbb{P}(\varepsilon) \right)^{1/q} = \left(\int_{\Omega} \left| \sum_{i=1}^n t_i \theta_i \varepsilon_i \right|^q d\mathbb{P}(\varepsilon) \right)^{1/q} \\ &\geq \left(\int_{Q_n} \left| \sum_{i=1}^n t_i \theta_i y_i \right|^q dy \right)^{1/q} = \|\langle \cdot, (t\theta) \rangle\|_{L^q(Q_n)}, \end{aligned}$$

where $t_i = \int_K |x_i| dx$ and $t\theta = (t_1 \theta_1, \dots, t_n \theta_n)$. Since $t_i \simeq 1$ for all $i = 1, \dots, n$, from (1.7) we readily see that

$$\|\langle \cdot, \theta \rangle\|_{L^q(K)} \geq \|\langle \cdot, (t\theta) \rangle\|_{L^q(Q_n)} \geq c \|\langle \cdot, \theta \rangle\|_{L^q(Q_n)}.$$

In view of (1.8), this observation and Theorem 1.1 show that, if K is unconditional, then a random K_N contains $Z_{\ln(N/n)}(Q_n)$:

Theorem 2.5 *Let $\beta \in (0, 1/2]$ and $\gamma > 1$. There exists an absolute constant $c > 0$ so that if*

$$(2.26) \quad N \geq N(\gamma, n) = c\gamma n,$$

and if $K_N = \text{conv}\{x_1, \dots, x_N\}$ is a random polytope spanned by N independent random points x_1, \dots, x_N uniformly distributed in an unconditional isotropic convex body K in \mathbb{R}^n , then we have

$$(2.27) \quad K_N \supseteq c_1 C(\alpha) = c_1 (\alpha B_2^n \cap B_\infty^n) \quad \text{for all } \alpha \leq c_2 \sqrt{\beta \ln(N/n)},$$

with probability greater than

$$(2.28) \quad 1 - \exp(-c_3 N^{1-\beta} n^\beta) - \mathbb{P}(\|\Gamma : \ell_2^n \rightarrow \ell_2^N\| \geq \gamma\sqrt{N}),$$

where $\Gamma : \ell_2^n \rightarrow \ell_2^N$ is the random operator $\Gamma(y) = (\langle x_1, y \rangle, \dots, \langle x_N, y \rangle)$ defined by the vertices x_1, \dots, x_N of K_N .

Next, we outline a direct proof of Theorem 2.5 (in which L_q -centroid bodies are not involved): For $k \in \mathbb{N}$ and $y \in \mathbb{R}^n$, define

$$(2.29) \quad \|y\|_{P(k)} := \sup \left\{ \sum_{i=1}^k \left(\sum_{j \in B_i} y_j^2 \right)^{1/2} : \bigcup_{i=1}^k B_i = [n], B_i \cap B_j = \emptyset (i \neq j) \right\},$$

where we write $[n]$ for the set $\{1, 2, \dots, n\}$. Montgomery–Smith has shown (see [19]) that: For any $y \in \mathbb{R}^n$ and $k \in \mathbb{N}$, one has

$$(2.30) \quad \mathbb{P} \left(\sum_{i=1}^n \varepsilon_i y_i \geq \lambda \|y\|_{P(k)} \right) \geq \left(\frac{1}{3} \right)^k (1 - 2\lambda^2)^{2k} \quad (0 \leq \lambda \leq 1/\sqrt{2}).$$

Also, for $y \in \mathbb{R}^n$, one has

$$(2.31) \quad \|y\|_{P(t^2)} \leq K_{1,2}(y, t) \leq \sqrt{2} \|y\|_{P(t^2)}$$

when $t^2 \in \mathbb{N}$, from where one concludes the following:

Lemma 2.6 *There exists a constant $c > 0$ such that, for all $y \in \mathbb{R}^n$ and any $t > 0$,*

$$(2.32) \quad \mathbb{P} \left(\sum_{i=1}^n \varepsilon_i y_i \geq \lambda K_{1,2}(y, t) \right) \geq e^{-\phi(\lambda)t^2},$$

where $\phi(\lambda) = 4 \ln(3(1 - 2\lambda^2)^{-2})$ for $0 < \lambda < 1/\sqrt{2}$.

P. Pivovarov [23] has recently obtained the following result: There exists an absolute constant $C \geq 1$ such that for any unconditional isotropic convex body K in \mathbb{R}^n , the spherical measure of the set of $\theta \in S^{n-1}$ such that

$$\mathbb{P}(|\langle x, \theta \rangle| \geq t) \geq \exp(-Ct^2)$$

whenever $C \leq t \leq \frac{\sqrt{n}}{C \ln n}$, is at least $1 - 2^{-n}$. The proof of the next Lemma follows more or less the same lines.

Lemma 2.7 *Let K be an isotropic unconditional convex body in \mathbb{R}^n . For every $\theta \in S^{n-1}$ and any $\alpha \geq 1$ we have*

$$(2.33) \quad \mathbb{P}_x(\langle x, \theta \rangle \geq h_{C(\alpha)}(\theta)) \geq c_1 e^{-c_2 \alpha^2}.$$

Proof. For $\theta = (\theta_i)_{i=1}^n \in \mathbb{S}^{n-1}$, $x \in K$ and $0 < s < 1/\sqrt{2}$ define the set

$$(2.34) \quad K_s(\theta) = \{x \in K : K_{1,2}(\theta, \alpha) \leq sK_{1,2}(x\theta, \alpha)\},$$

where by “ $x\theta$ ” we mean the vector with coordinates $x_i\theta_i$ and s is to be chosen. We have:

$$\begin{aligned} \mathbb{P}_x \left(\sum_{i=1}^n x_i \theta_i \geq h_{C(\alpha)}(\theta) \right) &= \mathbb{P}_x \left(\sum_{i=1}^n \varepsilon_i x_i \theta_i \geq h_{C(\alpha)}(\theta) \right) \\ &= \int_K \mathbb{P}_\varepsilon \left(\sum_{i=1}^n \varepsilon_i (x_i \theta_i) \geq h_{C(\alpha)}(\theta) \right) dx \\ &= \int_K \mathbb{P}_\varepsilon \left(\sum_{i=1}^n \varepsilon_i (x_i \theta_i) \geq K_{1,2}(\theta, \alpha) \right) dx \\ &\geq \int_{K_s(\theta)} \mathbb{P}_\varepsilon \left(\sum_{i=1}^n \varepsilon_i (x_i \theta_i) \geq sK_{1,2}(x\theta, \alpha) \right) dx \\ &\geq e^{-\phi(s)\alpha^2} |K_s(\theta)|, \end{aligned}$$

by Lemma 2.6.

Assume first that $m := \alpha^2$ is an integer and let B_1, B_2, \dots, B_m be a partition of the set $\{1, 2, \dots, n\}$ so that

$$(2.35) \quad K_{1,2}(\theta, \alpha) = \sum_{i=1}^m \left(\sum_{j \in B_i} |\theta_j|^2 \right)^{1/2} =: A.$$

Consider the seminorm

$$(2.36) \quad f(x) = \sum_{i=1}^m \left(\sum_{j \in B_i} |x_j \theta_j|^2 \right)^{1/2}.$$

Then, using the reverse Hölder inequality $c_1 \|f\|_{L^2(K)} \leq \|f\|_{L^1(K)}$ and the fact that $L_K \simeq 1$, we get

$$\begin{aligned} \int_K K_{1,2}(x\theta, \alpha) dx &\geq \int_K \sum_{i=1}^m \left(\sum_{j \in B_i} |x_j \theta_j|^2 \right)^{1/2} \\ &\geq c_1 \sum_{i=1}^m \left(\sum_{j \in B_i} |\theta_j|^2 \int_K |x_j|^2 \right)^{1/2} \\ &\geq cA. \end{aligned}$$

We now apply the Paley-Zygmund inequality to get

$$(2.37) \quad |K_s(\theta)| = \mathbb{P}_x (f > sA) \geq \frac{(\mathbb{E}|f|^2 - (sA)^2)^2}{\mathbb{E}[f^4]}.$$

Choosing $s = \frac{1}{2\sqrt{2}} \min\{c, 1\}$ we get

$$|K_s(\theta)| \geq \frac{cA^4}{\mathbb{E}[f^4]},$$

for a suitable new absolute constant $c > 0$. On the other hand, we can estimate $\mathbb{E}[f^4]$ from above, by the reverse Hölder inequality:

$$\begin{aligned} (\mathbb{E}[f^4])^{1/4} &\leq 4c\mathbb{E}|f| = 4c \sum_{i=1}^m \mathbb{E} \left(\sum_{j \in B_i} |x_j \theta_j|^2 \right)^{1/2} \\ &\leq 4cL_K \sum_{i=1}^m \left(\sum_{j \in B_i} |\theta_j|^2 \right)^{1/2} \leq 4cA. \end{aligned}$$

As a result, $|K_s(\theta)| \geq c$. Returning to the estimate

$$(2.38) \quad \mathbb{P}_x \left(\sum_{i=1}^n x_i \theta_i \geq h_{C(\alpha)}(\theta) \right) \geq e^{-\phi(s)\alpha^2} |K_s(\theta)|,$$

we get:

$$(2.39) \quad \mathbb{P}_x \left(\sum_{i=1}^n x_i \theta_i \geq h_{C(\alpha)}(\theta) \right) \geq ce^{-c\alpha^2}.$$

This proves the Lemma for $\alpha^2 \in \mathbb{N}$ and the result follows easily for every α . \square

Proof of Theorem 2.5. Now, using the procedure of the proof of Theorem 1.1 we complete the proof of Theorem 2.5. \square

Remark 2.8 Regarding the probability $\mathbb{P}(\|\Gamma : \ell_2^n \rightarrow \ell_2^N\| \geq \gamma\sqrt{N})$, in the unconditional case Aubrun has proved in [1] that for every $\rho > 1$ and $N \geq \rho n$, one has

$$(2.40) \quad \mathbb{P}(\|\Gamma : \ell_2^n \rightarrow \ell_2^N\| \geq c_1(\rho)\sqrt{N}) \leq \exp(-c_2(\rho)n^{1/5}).$$

In particular, one can find $c, C > 0$ so that, if $N \geq Cn$, then

$$(2.41) \quad \mathbb{P}(\|\Gamma : \ell_2^n \rightarrow \ell_2^N\| \geq C\sqrt{N}) \leq \exp(-cn^{1/5}).$$

This allows us to use Theorem 2.5 with a probability estimate $1 - \exp(-cn^e)$ for values of N which are proportional to n .

3 Weakly sandwiching K_N

We proceed to the question whether the inclusion given by Theorem 1.1 is sharp. It was already mentioned in the Introduction that we cannot expect a reverse inclusion of the form $K_N \subseteq c_4 Z_q(K)$ with probability close to 1, unless if q is of the order of n . To see this, observe that, for any $\alpha > 0$,

$$\begin{aligned} \mathbb{P}(K_N \subseteq \alpha Z_q(K)) &= \mathbb{P}(x_1, x_2, \dots, x_N \in \alpha Z_q(K)) \\ &= \left(\mathbb{P}(x \in \alpha Z_q(K)) \right)^N \\ &\leq |\alpha Z_q(K)|^N. \end{aligned}$$

It was proved in [20] that, for every $q \leq n$, the volume of $Z_q(K)$ is bounded by $(c\sqrt{q/n}L_K)^n$. This leads immediately to the estimate

$$(3.1) \quad \mathbb{P}(K_N \subseteq \alpha Z_q(K)) \leq (c\alpha\sqrt{q/n}L_K)^{nN},$$

where $c > 0$ is an absolute constant. Assume that K has bounded isotropic constant and we want to keep $\alpha \simeq 1$. Then, (3.1) shows that, independently from the value of N , we have to choose q of the order of n so that it might be possible to show that $\mathbb{P}(K_N \subseteq \alpha Z_q(K))$ is really close to 1. Actually, if $q \sim n$ then this is always the case, because $Z_n(K) \supseteq cK$.

Lemma 3.1 *Let K be a convex body of volume 1 in \mathbb{R}^n and let $N > n$. Fix $\alpha > 1$. Then, for every $\theta \in S^{n-1}$ one has*

$$(3.2) \quad \mathbb{P}(h_{K_N}(\theta) \geq \alpha h_{Z_q(K)}(\theta)) \leq N\alpha^{-q}.$$

Proof. Markov's inequality shows that

$$(3.3) \quad \mathbb{P}(\alpha, \theta) := \mathbb{P}(x \in K : |\langle x, \theta \rangle| \geq \alpha \|\langle \cdot, \theta \rangle\|_q) \leq \alpha^{-q}.$$

Then,

$$\mathbb{P}(h_{K_N}(\theta) \geq \alpha h_{Z_q(K)}(\theta)) = \mathbb{P}(\max_{j \leq N} |\langle x_j, \theta \rangle| \geq \alpha \|\langle \cdot, \theta \rangle\|_q) \leq N\mathbb{P}(\alpha, \theta)$$

and the result follows. \square

Lemma 3.2 *Let K be a convex body of volume 1 in \mathbb{R}^n and let $N > n$. For every $\alpha > 1$ one has*

$$(3.4) \quad \mathbb{E} [\sigma(\theta : (h_{K_N}(\theta) \geq \alpha h_{Z_q(K)}(\theta)))] \leq N\alpha^{-q}.$$

Proof. Immediate: observe that

$$\mathbb{E} [\sigma(\theta : (h_{K_N}(\theta) \geq \alpha h_{Z_q(K)}(\theta)))] = \int_{S^{n-1}} \mathbb{P}(h_{K_N}(\theta) \geq \alpha h_{Z_q(K)}(\theta)) d\sigma(\theta)$$

by Fubini's theorem. \square

The estimate of Lemma 3.2 is already enough to show that if $q \geq c \ln N$ then, on the average, $h_{K_N}(\theta) \leq ch_{Z_q(K)}(\theta)$ with probability greater than $1 - N^{-c}$. In particular, the mean width of a random K_N is bounded by the mean width of $Z_{\ln(N/n)}(K)$:

Proposition 3.3 *Let K be an isotropic convex body in \mathbb{R}^n . If $q \geq 2 \ln N$ then*

$$(3.5) \quad \mathbb{E} [w(K_N)] \leq cw(Z_q(K)),$$

where $c > 0$ is an absolute constant.

Proof. We write

$$(3.6) \quad w(K_N) \leq \int_{A_N} h_{K_N}(\theta) d\sigma(\theta) + c\sigma(A_N^c)nL_K,$$

where $A_N = \{\theta : h_{K_N}(\theta) \leq \alpha h_{Z_q(K)}(\theta)\}$. Then,

$$(3.7) \quad w(K_N) \leq \alpha \int_{A_N} h_{Z_q(K)}(\theta) d\sigma(\theta) + c\sigma(A_N^c)nL_K,$$

and hence, by Lemma 3.2,

$$(3.8) \quad \mathbb{E} w(K_N) \leq \alpha w(Z_q(K)) + cNn\alpha^{-q}L_K.$$

Since $w(Z_q(K)) \geq w(Z_2(K)) = L_K$, we get

$$(3.9) \quad \mathbb{E} w(K_N) \leq (\alpha + cNn\alpha^{-q})w(Z_q(K)).$$

The result follows if we choose $\alpha = e$. \square

3.1 Volume radius of K_N

Next, we discuss the volume radius of K_N . A lower bound follows by comparison with the Euclidean ball. It was proved in [12, Lemma 3.3] that if K is a convex body in \mathbb{R}^n with volume 1, then

$$(3.10) \quad \mathbb{P}(|K_N| \geq t) \geq \mathbb{P}(|\overline{B}_2^n|_N \geq t)$$

for every $t > 0$. Therefore, it is enough to consider the case of B_2^n . In [10] it is shown that there exist $c_1 > 1$ and $c_2 > 0$ such that if $N \geq c_1 n$ and x_1, \dots, x_N are independent random points uniformly distributed in \overline{B}_2^n , then

$$(3.11) \quad [\overline{B}_2^n]_N \supseteq c_2 \min \left\{ \frac{\sqrt{\ln(2N/n)}}{\sqrt{n}}, 1 \right\} \overline{B}_2^n$$

with probability greater than $1 - \exp(-n)$. It follows that if $N \geq c_1 n$ then, with probability greater than $1 - \exp(-n)$ we have

$$(3.12) \quad |K_N|^{1/n} \geq c_2 \min \left\{ \frac{\sqrt{\ln(2N/n)}}{\sqrt{n}}, 1 \right\},$$

where $c_1 > 1$ and $c_2 > 0$ are absolute constants.

The case $n < N < c_1 n$ was studied in [7] where it was proved that (3.11) continues to hold true with probability greater than $1 - \exp(-cn/\ln n)$, where $c > 0$ is an absolute constant. Combining this fact with (3.10), we see that (3.12) is valid for all $N > n$.

We now pass to the upper bound; Proposition 3.3, combined with Urysohn's inequality, yields the following:

Proposition 3.4 *Let K be an isotropic convex body in \mathbb{R}^n . If $N > n$ and $q \geq 2 \ln N$, then*

$$(3.13) \quad \mathbb{E}(K, N) \leq \frac{c_1 \mathbb{E}[w(K_N)]}{\sqrt{n}} \leq \frac{c_2 w(Z_q(K))}{\sqrt{n}},$$

where $c_1, c_2 > 0$ are absolute constants.

Proposition 3.4 reduces, in a sense, the question to that of giving upper bounds for $w(Z_q(K))$. It is proved in [20] that, if $q = \ln N \leq \sqrt{n}$ then $w(Z_q(K)) \leq c\sqrt{q}L_K$. It follows that

$$(3.14) \quad \mathbb{E}(K, N) \leq c \frac{\sqrt{\ln(N/n)}L_K}{\sqrt{n}},$$

which is the conjectured estimate for $N \leq e^{\sqrt{n}}$. For $q = \ln N > \sqrt{n}$ we know that $w(Z_q(K)) \leq \frac{qL_K}{\sqrt{n}}$ since $Z_q(K) \subseteq (q/\sqrt{n})Z_{\sqrt{n}}(K)$. This is most probably a non-optimal bound.

However, we can further exploit the simple estimate of Lemma 3.1 to obtain a sharp estimate for larger values of N . We will make use of the following facts:

Fact 1. Let A be a symmetric convex body in \mathbb{R}^n . For any $1 \leq q < n$, set

$$(3.15) \quad w_{-q}(A) = \left(\int_{S^{n-1}} \frac{1}{h_A^q(\theta)} d\sigma(\theta) \right)^{-1/q}.$$

An application of Hölder's inequality shows that

$$(3.16) \quad \left(\frac{|A^\circ|}{|B_2^n|} \right)^{1/n} = \left(\int_{S^{n-1}} \frac{1}{h_A^n(\theta)} d\sigma(\theta) \right)^{1/n} \geq \left(\int_{S^{n-1}} \frac{1}{h_A^q(\theta)} d\sigma(\theta) \right)^{1/q} = \frac{1}{w_{-q}(A)}.$$

From the Blaschke–Santaló inequality, it follows that

$$(3.17) \quad |A|^{1/n} \leq |B_2^n|^{1/n} w_{-q}(A) \leq \frac{c_1 w_{-q}(A)}{\sqrt{n}}.$$

Fact 2. A recent result of G. Paouris (see [21, Proposition 5.4]) shows that if A is an isotropic convex body in \mathbb{R}^n then, for any $1 \leq q < n/2$,

$$(3.18) \quad w_{-q}(Z_q(A)) \simeq \frac{\sqrt{q}}{\sqrt{n}} I_{-q}(A)$$

where

$$(3.19) \quad I_p(A) = \left(\int_A \|x\|_2^p dx \right)^{1/p}, \quad p > -n.$$

Fact 3. Let K be an isotropic convex body in \mathbb{R}^n , let $N > n^2$ and $q = 2 \ln(2N)$. We write

$$\begin{aligned} [w_{-q/2}(Z_q(K))]^{-q} &= \left(\int_{S^{n-1}} \frac{1}{h_{Z_q(K)}^{q/2}(\theta)} d\sigma(\theta) \right)^2 \\ &\leq \left(\int_{S^{n-1}} \frac{1}{h_{K_N}^q(\theta)} d\sigma(\theta) \right) \left(\int_{S^{n-1}} \frac{h_{K_N}^q(\theta)}{h_{Z_q(K)}^q(\theta)} d\sigma(\theta) \right). \end{aligned}$$

Observe that $K_N \subseteq K \subseteq (n+1)L_K$ and $Z_q(K) \supseteq Z_2(K) \supseteq L_K B_2^n$, and hence, $h_{K_N}(\theta) \leq (n+1)h_{Z_q(K)}(\theta)$ for all $\theta \in S^{n-1}$. Therefore,

$$(3.20) \quad \int_{S^{n-1}} \frac{h_{K_N}^q(\theta)}{h_{Z_q(K)}^q(\theta)} d\sigma(\theta) = \int_0^{n+1} qt^{q-1} [\sigma(\theta : h_{K_N}(\theta) \geq th_{Z_q(K)}(\theta))] dt.$$

Fact 4. Taking expectations in (3.20) and using Lemma 3.2, we see that, for every $a > 1$,

$$\begin{aligned} \mathbb{E} \left[\int_{S^{n-1}} \frac{h_{K_N}^q(\theta)}{h_{Z_q(K)}^q(\theta)} d\sigma(\theta) \right] &\leq a^q + \int_a^{n+1} qt^{q-1} Nt^{-q} dt \\ &= a^q + qN \ln \left(\frac{n+1}{a} \right). \end{aligned}$$

Choosing $a = 2e$ and using the fact that $e^q = (2N)^2$ by the choice of q , we see that

$$(3.21) \quad \mathbb{E} \left[\int_{S^{n-1}} \frac{h_{K_N}^q(\theta)}{h_{Z_q(K)}^q(\theta)} d\sigma(\theta) \right] \leq c_2^q$$

where $c_2 > 0$ is an absolute constant. Then, Markov's inequality implies that

$$(3.22) \quad \int_{S^{n-1}} \frac{h_{K_N}^q(\theta)}{h_{Z_q(K)}^q(\theta)} d\sigma(\theta) \leq (c_2 e)^q$$

with probability greater than $1 - e^{-q}$. Going back to Fact 3, we conclude that $[w_{-q/2}(Z_q(K))]^{-q} \leq c_3^q [w_{-q}(K_N)]^{-q}$, i.e.

$$(3.23) \quad w_{-q}(K_N) \leq c_4 w_{-q/2}(Z_q(K))$$

with probability greater than $1 - e^{-q}$.

Proof of Theorem 1.3. Assume that K_N satisfies (3.23) and set $S_N = K_N - K_N$. From Fact 1 we have

$$(3.24) \quad |K_N|^{1/n} \leq |S_N|^{1/n} \leq \frac{c_1}{\sqrt{n}} w_{-q}(S_N) = \frac{2c_1}{\sqrt{n}} w_{-q}(K_N).$$

Now, Fact 4 shows that

$$(3.25) \quad |K_N|^{1/n} \leq \frac{c_5}{\sqrt{n}} w_{-q/2}(Z_q(K))$$

with probability greater than $1 - e^{-q}$. Since $Z_q(K) \subseteq cZ_{q/2}(K)$, using Fact 2 we write

$$(3.26) \quad w_{-q/2}(Z_q(K)) \leq c_6 w_{-q/2}(Z_{q/2}(K)) \leq \frac{c_7 \sqrt{q}}{\sqrt{n}} I_{-q/2}(K).$$

Since K is isotropic, we have $I_{-q/2}(K) \leq I_2(K) = \sqrt{n} L_K$, which implies

$$(3.27) \quad w_{-q/2}(Z_q(K)) \leq c_7 \sqrt{q} L_K.$$

Putting everything together, we have

$$(3.28) \quad |K_N|^{1/n} \leq \frac{c\sqrt{q}}{\sqrt{n}} L_K \simeq \frac{\sqrt{\ln(N/n)} L_K}{\sqrt{n}},$$

with probability greater than $1 - e^{-q} \geq 1 - \frac{1}{N}$. This completes the proof. \square

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